

Gene flow biases population genetic inference of recombination rate

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1 Abstract

2 Accurate estimates of the rate of recombination are key to understanding a host of
3 evolutionary processes as well as the evolution of recombination rate itself. Model-
4 based population genetic methods that infer recombination rates from patterns of
5 linkage disequilibrium (LD) in the genome have become a popular method to estimate
6 rates of recombination. However, these LD-based methods make a variety of simplifying
7 assumptions about the populations of interest that are often not met in natural
8 populations. One such assumption is the absence of gene flow from other populations.
9 Here, we use forward-time population genetic simulations of isolation-with-migration
10 scenarios to explore how gene flow affects the accuracy of LD-based estimators of
11 recombination rate. We find that moderate levels of gene flow can result in either the
12 overestimation or underestimation of recombination rates by up to 20-50% depending
13 on the timing of divergence. We also find that these biases can affect the detection of
14 interpopulation differences in recombination rate, causing both false positive and false
15 negatives depending on the scenario. We discuss future possibilities for mitigating
16 these biases and recommend that investigators exercise caution and confirm that their
17 study populations meet assumptions before deploying these methods.

18 Introduction

19 Recombination rate, the number of crossovers per unit genome per generation, plays a
20 key role in shaping evolutionary processes and diversity in the genome. For example,
21 through the action of linked selection, local rates of recombination are a chief
22 determinant of patterns of genetic diversity throughout the genome (Begun & Aquadro,

23 1992; Burri, 2017; Cutter, 2019; Haddrill et al., 2014; Korunes et al., 2021). Genome-wide
24 rates of recombination also modulate diverse processes such as adaptation, speciation,
25 and introgression (Dapper & Payseur, 2017; Samuk et al., 2017; Schumer et al., 2018;
26 Stapley et al., 2017). There is also a growing appreciation that recombination rate is
27 itself a trait that varies and evolves (Dumont & Payseur, 2008; Hunter et al., 2016;
28 Johnston et al., 2016; Ritz et al., 2017; Samuk et al., 2020; Stapley et al., 2017).
29 Accordingly, there has been great interest in efficient and accurate methods for
30 estimating recombination rates.

31 Current methods for estimating recombination rates fall into two broad classes of
32 methods: direct and indirect (Peñalba & Wolf, 2020). Of the direct measures, the three
33 most popular approaches are linkage mapping, gamete sequencing, and cytological
34 methods. With classical linkage mapping, map distances between genetic markers are
35 measured by quantifying recombinant markers in the context of a genetic cross or
36 pedigree (Broman, 2010; Rastas, 2017). The resolution of this approach is limited only by
37 marker density and the sample size of individuals, but larger sample sizes can be
38 grueling to carry out in the laboratory or unavailable in some populations. Further,
39 identifying suitable diagnostic mapping markers can be limiting in some cases (e.g. in a
40 highly homozygous population, Broman, 2010). Direct sequencing of pools of
41 recombinant gamete genomes from single individuals using long/linked read
42 sequencing is a newer approach that alleviates many of the issues of traditional
43 mapping, but still requires differentiated markers to score crossover events between
44 homologous chromosomes (Dréau et al., 2019; Rommel Fuentes et al., 2019; Xu et al.,
45 2020). Cytological methods bypass this requirement by directly visualizing

46 recombination-associated protein complexes in cell populations undergoing meiosis
47 (Peterson et al., 2019; Peterson & Payseur, 2021). However, cytological methods are
48 limited by the spatial resolution at which such visualization can occur (e.g. the
49 resolution of immunostained gamete karyotypes, Peterson et al., 2019).

50 Because all direct methods of measuring recombination rates are fairly laborious, there
51 has been increased interest in indirect measures of recombination rate that leverage
52 readily available population genetic data. Chief among these are model-based methods
53 that infer rates of recombination from patterns of linkage disequilibrium (LD), (Auton &
54 McVean, 2007; Chan, Song, et al., 2012; Kamm et al., 2016; Spence & Song, 2019). These
55 methods attempt to estimate recombination rates by statistically fitting recombination
56 rates (derived from population genetic models/simulations) to observed patterns of LD.
57 Rather than inferring recombination rate directly, LD-based estimators infer a
58 *population scaled recombination rate*, $\rho = 4N_e r$, where N_e is the effective population size
59 and r is the theoretical per-generation recombination rate. LD-based methods are
60 attractive because they (1) generally only require population-scale genomic data and (2)
61 are very fast, often only requiring several computational hours or less (Spence & Song,
62 2019) and (3) are informative of time-averaged population historical recombination rates
63 (Gil McVean & Auton, 2007). Accordingly, LD-based estimates of recombination rates
64 have become extremely popular, and now vastly outnumber direct measures in the
65 literature (Peñalba & Wolf, 2020; Stapley et al., 2017). These methods have also begun to
66 be used to perform interpopulation comparisons of recombination rates (Peñalba &
67 Wolf, 2020; Stapley et al., 2017).

68 Like all models, LD-based estimators of recombination rate make a variety of
69 simplifying assumptions about the populations of interest. For one, they generally
70 assume that the populations/loci of interest are evolving largely neutrally and have
71 reached population genetic equilibrium in a number of ways (Stumpf & McVean, 2003).
72 In particular, most methods assume that the populations being studied have reached an
73 equilibrium between recombination and population scaled mutation, such that LD
74 accurately reflects patterns of recombination rate (G. McVean, 2007). Further, it is
75 generally assumed that any form of selection that might distort patterns of LD (e.g.
76 sweeps) has not recently occurred (Chan, Song, et al., 2012). Finally, these methods make
77 the general assumption that demographic processes that distort genome-wide patterns
78 of LD, such as population size changes, have not occurred (recall that ρ is directly
79 dependent on N_e , Auton & McVean, 2007).

80 While some of these assumptions may be robust to violation, work has shown that some
81 violations can result in biased estimates. For example, (Dapper & Payseur, 2018) showed
82 that recombination estimates from LDhat (Gil McVean & Auton, 2007) are highly
83 sensitive to changes in population size. This can be ameliorated in some cases by
84 incorporating known changes in population size into the estimation procedure, such as
85 implemented in the software pyrho (Spence & Song, 2019).

86 Along with changes in population size and selection, another process that can greatly
87 alter patterns of LD is gene flow. Gene flow and subsequent admixture between
88 diverged populations can have complex effects on patterns of LD within each
89 population (Nei & Li, 1973; Ohta, 1982). These effects range from large and genomically

90 variable increases in LD due to segregation of divergent haplotypes, to genome-wide
91 decreases in LD as populations become coupled and increase local N_e (Nei & Li, 1973;
92 Ohta, 1982). While it is now widely accepted that gene flow is commonplace in natural
93 populations (Barton, 2008; Mallet, 2005; Suvorov et al., 2021; Waples & Gaggiotti, 2006),
94 and there has not been a systematic study of the effects of gene flow on LD-based
95 measures of recombination. Further, it remains unclear how gene flow (or any other
96 violation of assumptions) impacts our ability to detect differences in recombination rate
97 *between* (as opposed to *within*) populations using LD-based methods.

98 Here, we address these issues using forward-time population genetic simulations. We
99 attempt to answer two specific questions. First, how does gene flow between
100 populations affect the precision and accuracy of LD-based estimates of recombination
101 rate within populations? Secondly, how does gene flow affect our ability to detect
102 evolved differences in recombination rate between populations? Our primary goal is to
103 answer these questions in the context of a core set of realistic demographic scenarios,
104 and not perform an exhaustive exploration of parameter space. Overall we hope to help
105 investigators understand key sources of bias in LD-based estimates of recombination
106 rate in natural populations and highlight areas of future development.

107 Methods

108 Code availability

109 All scripts used in the analyses described below are available as a repository on Github
110 (http://github.com/ksamuk/LD_recomb).

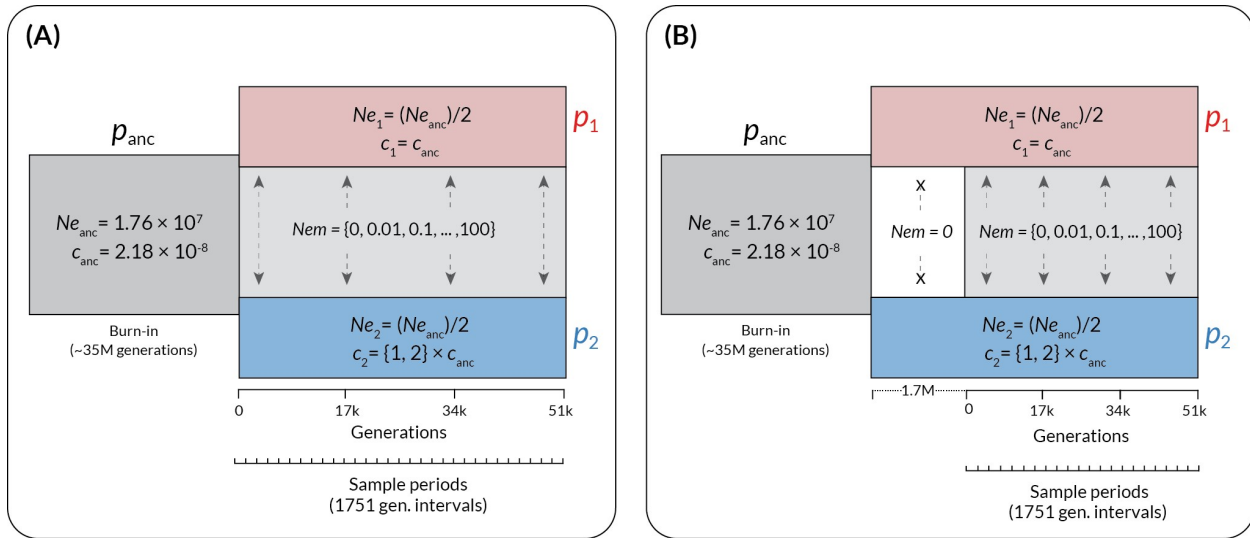
111 Forward time simulations with SLiM

112 To explore how the timing and amount of gene flow affect estimates of recombination
113 rate, we performed forward-time population genetic simulations using SLiM version 3.3
114 (Haller & Messer, 2019). The basic form of all the simulations was an isolation-with-
115 migration scenario: a single ancestral population diverges into two subpopulations with
116 a static amount of bidirectional gene flow (Figure 1). Populations were composed of
117 diploid individuals with 100kb genomes arranged in a single chromosome. We used
118 genome-wide average estimates of effective population size, mutation rate, and
119 empirical recombination rate from natural populations of *Drosophila melanogaster*
120 (Adrion, Cole, et al., 2020): Mutation rate = 5.49×10^{-9} (Li & Stephan, 2006); Recombination
121 rate = 2.23×10^{-8} , (average of chromosome 2R, Comeron et al., 2012); $N_e = 1.72M$ (Li &
122 Stephan, 2006). Recombination and mutation rates were conservatively modeled as
123 uniform across the 100kb genome. Following standard practice for forward-time
124 simulations, all simulations were run with an *in silico* population size of $N=1000$, and
125 simulated mutation and recombination rates scaled by a factor of N_e/N as per the SLiM
126 manual (Haller & Messer, 2019). Note that generation times are also subject to scaling,
127 and for simplicity, we will refer to all generations in terms of back-transformed actual
128 generations rather than SLiM generations (1 SLiM generation ≈ 1751 actual generations
129 with our scaling factor).

130 Parameter space

131 To explore how variation in gene flow affects estimates of recombination, we varied the
132 amount of gene flow over five orders of magnitude: 0, 0.01, 0.1, 1, 10, 100, in standard
133 units of $N_e m$ (the product of the effective population size and the migration rate). These

134 values were chosen to encompass total isolation ($N_e m = 0$), limited gene flow ($N_e m =$
135 $0.01-0.1$), moderate gene flow in interconnected metapopulations ($N_e m = 1-10$, Morjan &
136 Rieseberg, 2004; Waples & Gaggiotti, 2006), and a scenario of a nascent hybrid swarm
137 ($N_e m = 100$). We also varied the timing of the onset of gene flow, with gene flow
138 beginning either immediately after divergence or after a period of isolation. We
139 performed preliminary simulations to determine a period of isolation ($\sim 1.7M$
140 generations in our case) that produced levels of genomic divergence (Figure S1) similar
141 to those observed in natural population pairs that exhibit genome-wide genetic
142 divergence but still actively exchange genes ($F_{ST} \sim 0.4$, Morjan & Rieseberg, 2004; Roux
143 et al., 2016). Finally, to explore how gene flow impacts the detection of population
144 *differences* in recombination rate, we modeled scenarios where recombination rate either
145 remains constant in both subpopulations or instantaneously increases by a factor of 2 at
146 the time of divergence in one of the two subpopulations (always subpopulation two).
147 This magnitude of this difference is well within the range of variation in recombination
148 rate reported for a wide variety of species (Stapley et al., 2017). In biological terms, an
149 instantaneous increase in population recombination rate could be readily mediated by
150 an environmental change (e.g. temperature, Lloyd et al., 2018) or via a change in mating
151 system (Brandvain & Wright, 2016). We note that this instantaneous change is a “best
152 case” scenario for detecting interpopulation differences in recombination rate, and thus
153 any loss of power to detect differences in recombination that occurs due to gene flow
154 will be conservative.



155 **Figure 1** | The structure of the forward-time simulations performed in SLiM. Time in back-transformed
 156 generations is shown along the x-axis, and the populations in existence at a given time are shown as
 157 rectangles. p_{anc} = the ancestral population, p_1 = the subpopulation with unchanged recombination rate, and
 158 p_2 = the subpopulation with increased recombination rate (if applicable). Effective population sizes (N_e)
 159 and recombination rates (c , in units of cM/Mb) are shown for each population, with the values for the
 160 subpopulations shown relative to the ancestral value. Variable elements of the simulation are shown in
 161 braces. Time in generations post-divergence is indicated below the plots, with the pre-contact isolation
 162 period in (B) shown as a dotted line preceding the main axis. Sample periods indicate intervals at which
 163 genotypes were output for analysis.

164 Details of demographic events

165 Each simulation began with a single population of size $N_{e_{anc}}$, which evolved for a 35M
 166 generation burn-in period (following the general practice of a 10-20 N_e burn-in period,
 167 Haller & Messer, 2019). This initial period was followed by divergence into two
 168 subpopulations, each with size $N_{e_{anc}}/2$. Gene flow (for cases where $N_{em} > 0$) began
 169 immediately at the time of divergence or after a 1.7M generation period of isolation and
 170 was symmetrical in magnitude and bidirectional. Changes in recombination rate

171 occurred at the time of divergence and instantaneously applied to all individuals in
172 subpopulation two only.

173 Starting at the time of divergence and thereafter in intervals of 1751 generations, we
174 collected a random sample of 25 individuals (a total of 50 haploid genomes) from each
175 population and saved their complete genotypic at all sites in VCF format. We stopped
176 the simulations after 51 000 generations. Each parameter combination was replicated
177 100 times, for a total number of $\sim n=48\ 000$ population samples.

178 Estimation of recombination rate using pyrho

179 While there are a variety of LD-based estimators of recombination rate, we elected to
180 use pyrho (Spence & Song, 2019) for estimation in this study. It shares its statistical
181 foundation with the most widely used LD-based estimators (LD-hat & LD-helmet; Chan,
182 Jenkins, et al., 2012; Gil McVean & Auton, 2007) while also having the ability to account
183 for changes in effective population size such as we are modeling here (Spence & Song,
184 2019). As such, any estimation biases caused by gene flow will likely affect those
185 approaches as least as much they affect pyrho. Direct comparisons with other methods
186 are complicated by the fact that pyrho is the only model-based method that adequately
187 accounts for changes in effective population (ReLERNN being a possible exception
188 Adrion, Galloway, et al., 2020)

189 We followed the recommended practices for inferring recombination rate using pyrho
190 (<https://github.com/popgenmethods/pyrho>). We parameterized the initial lookup tables
191 using the effective population size and mutation rates used in the simulations (*unscaled*

192 in this case). To account for changes in effective population size, we created lookup
193 tables that accounted for a change of $N_e/2$ (1.72M to 8.6M) in time steps of 1751
194 generations in the past. This allowed us to have an appropriately timed lookup table for
195 each step of the simulation. We used the built-in methods to infer the hyperparameters
196 of window size (best fit 100) and block penalty (best fit 1000). Using this baseline, we
197 inferred recombination rates using the genotype data (VCF format) from both
198 subpopulations at each time point, for a total of ~96 000 pyrho fits. All computation was
199 performed using the Duke University Computing Cluster, running CentOS Version 8.

200 Statistical analyses

201 We performed all data processing and visualization using the tools of the tidyverse
202 package in R 4.0.3 (R Core Team, 2018; Wickham, 2017). To examine how gene flow
203 between populations affects the accuracy of LD-based estimates of recombination, and
204 the context of the various factors explored in our simulations, we performed an analysis
205 of variance using a linear mixed model with Gaussian errors fitted via the `lmer()`
206 function from the `lme4` package (Bates et al., 2007). This model had the following form:
207 $\text{Recombination rate} = (1|\text{simulation replicate}) + (1|\text{simulation generation}) + \text{gene flow}$
208 $\text{magnitude} + \text{recombination rate change}$, where $(1|[\text{factor}])$ denotes a random intercept
209 and “:” denotes an interaction. All variables were standardized (mean-centered and
210 scaled by standard deviation) prior to analysis. To simplify interpretation, we fitted
211 separate models for the continuous gene flow and secondary contact scenarios.

212 Results

213 Inference when recombination rate is identical between populations

214 When the recombination rate remained constant between diverging populations, we
215 found that gene flow introduced two types of systematic biases in estimates of
216 recombination rate within populations (Figure 2A). These effects began when $N_e m \geq 1$
217 in both the continuous gene flow and secondary contact models. First, in the model of
218 continuous gene flow, when $N_e m \geq 1$, we observed a systematic increase (overestimate)
219 in estimated rates of recombination in both populations (Figure 2A, 2B, top row, $N_e m =$
220 1-100). This increase was statistically significant (Type III Wald chi-square = 5090.07, $p <$
221 2.0×10^{-16} ; coefficient for gene flow = 0.63–0.67 (95% CI), $t(19495) = 71.34$, $p < 0.001$). When
222 the migration rate was moderate to high ($N_e m$ 10-100), the recombination rate was
223 overestimated by approximately 10-20% (Figure 2B). This effect is consistent with
224 migration causing the populations to become coupled, behaving as a single population
225 with a larger N_e and thus inflating the population-scaled estimate of recombination
226 rate.

227 In contrast to the continuous gene flow case, under a model of secondary contact, there
228 was a marked systematic decrease (underestimate) of recombination rates, which also
229 became visible when $N_e m \geq 1$ (Figure 2A, 2B, bottom row, $N_e m = 1-100$). This decrease
230 was statistically significant (Type III Wald chi-square = 1512, $p < 2.2 \times 10^{-16}$; coefficient for
231 gene flow = -(0.54–0.49) (95% CI), $t(31846) = -38.88$, $p < 0.001$). The magnitude of this
232 decrease was substantial: on average, populations experiencing $N_e m = 1$ had
233 recombination rates about 20% lower than expected, with this increasing to 50% when

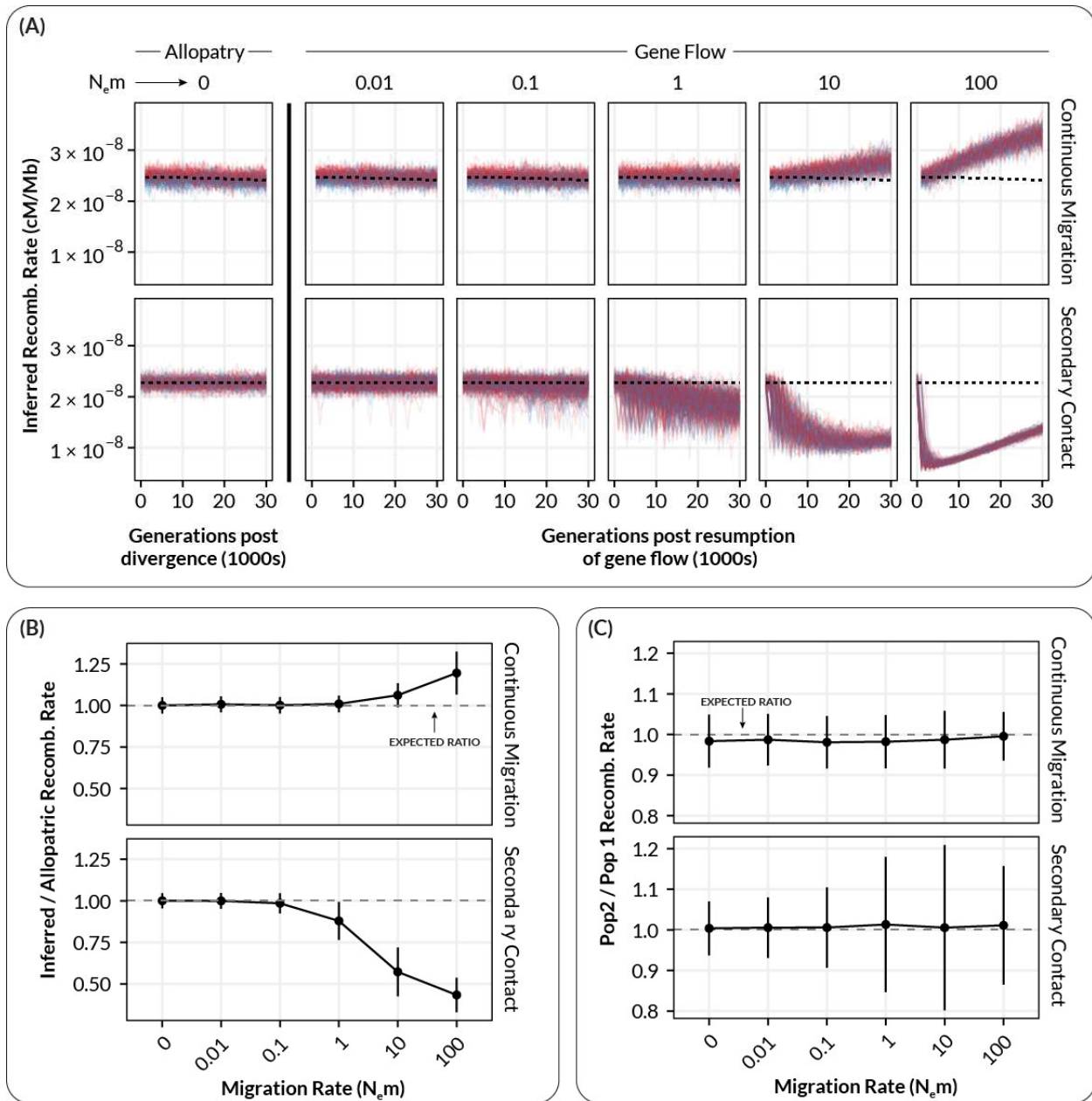
234 $N_e m = 10$ or higher (Figure 2A, Figure 2B). This decrease was accompanied by a
235 statistically significant increase in the *variance* of recombination rate estimates,
236 especially for $N_e m = 1-10$ compared to $N_e m < 1$ (Fig 2A, bottom row; F-test for
237 equivalency of variance, $F(10429,13860) = 0.20863$, $p < 2.2 \times 10^{-16}$). A systematic increase in
238 the mean and variance of LD within populations is consistent with allele frequency
239 differences between populations manifesting as migration-associated LD, and deflating
240 estimates of recombination rate. When gene flow was very high, there was a visible
241 recovery of estimated recombination rates (Figure 2A, bottom row, $N_e m = 100$),
242 presumably due to migration homogenizing allele frequencies and increased effective
243 population sizes increasing the rate at which recombination breaks down migration-
244 associated LD.

245 When comparing recombination rates between p_1 and p_2 , the “coupling” bias observed
246 in the continuous migration scenario did not appear to systematically affect the *ratio* of
247 recombination rate between the two populations (Figure 2C, Continuous Migration).
248 However, in keeping with the previous result, migration-associated LD in the secondary
249 contact model appeared to greatly increase the variance in the ratio of recombination
250 rates between populations when $N_e m \geq 1$ (Figure 2C, Secondary Contact).

251 Inference when recombination rate differs between populations

252 When recombination rates diverged between populations, we also observed the two
253 forms of bias described above (Figure 3). The estimates from the continuous gene flow
254 scenario exhibited a statistically significant increase (Type III Wald chi-square =
255 8936.44 , $p < 2.2 \times 10^{-16}$; coefficient for gene flow = $0.65-0.67$ (95% CI), $t(19495) = 94.53$, $p <$

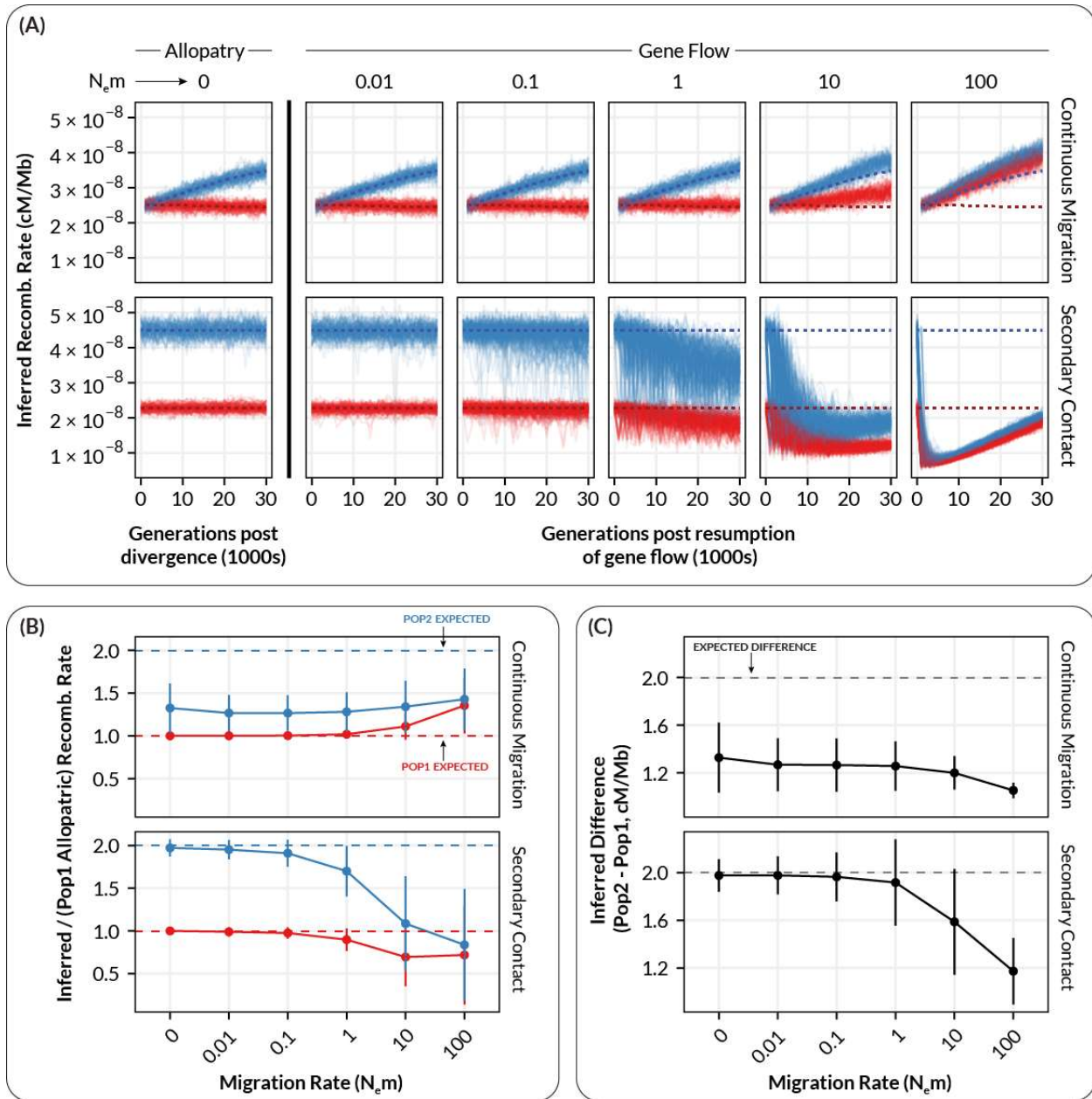
256 0.001) whereas estimates from the secondary contact model exhibited a statistically
257 significant decrease (Type III Wald chi-square = 1512, $p < 2.0 \times 10^{-16}$; coefficient for gene
258 flow = $-(0.27-0.22)$ (95% CI), $t(34505) = -23.22$, $p < 0.001$). However, the results differed
259 from simulations with constant recombination rates in a number of important ways.
260 First, there was a clear difference between the continuous migration and secondary
261 contact models in the overall trajectory in the population-specific estimates of
262 recombination rate (Figure 3A). In the continuous gene flow models, there was an
263 overall positive trend for the estimates of recombination rate in p_2 even in the absence
264 of gene flow (Figure 3A, continuous migration). This was presumably caused by a lag in
265 the establishment of equilibrium levels of LD within p_2 that reflect the new
266 recombination rate (which spontaneously changed at the time of divergence). This lag
267 resulted in the recombination rate in p_2 being consistently underestimated (because it
268 had not reached its new equilibrium), in addition to the coupling effect observed
269 previously (Figure 3B and 3C, continuous migration). In the case of the secondary
270 contact model, we did not observe the same positive trend for recombination rate
271 estimates in p_2 , likely because the isolation period (1.7M generations) was sufficiently
272 long enough for p_2 to establish an equilibrium level of LD prior to secondary contact
273 (Figure 3A, Secondary Contact).



274 **Figure 2** | The relationship between inferred recombination rate and the migration rate in simulated
 275 populations where recombination rate remains constant in both subpopulations. (A) Inferred
 276 recombination rates for individual simulations at varying levels of migration. Each plot shows inferred
 277 rates for simulation replicates (transparent lines) of population 1 (red, unchanged recombination) and
 278 population 2 (blue, increased recombination) for a single migration rate. Dashed lines show the expected
 279 inferred value in the absence of gene flow (inferred from $N_e m = 0$). (B) Summarized inferred recombination
 280 rates (y-axis) for each level of migration (x-axis) from the simulations in A. Points are mean values and

281 error bars depict standard deviations (summarized across all generations). Dashed lines show the
282 expected inferred value in the absence of gene flow for each population (i.e. the mean value for $N_e m = 0$).
283 (C) The inferred *difference* in recombination rate between population 1 and population 2 ($p_2 - p_1$) as a
284 function of migration rate. Points and errors bars are as in B.

285 In keeping with the scenario with constant recombination rates, starting at $N_e m \sim 1$,
286 migration-associated LD resulted in the systematic underestimation and increase in
287 variance for estimated recombination rates within both p_1 and p_2 (Figure 3B, Secondary
288 Contact; Type III Wald chi-square = 538.97, $p < 2.0 \times 10^{-16}$; coefficient for gene flow = -
289 (0.27-22) (95% CI), $t(34505) = -23.22$, $p < 0.001$; F-test for equivalency of variance,
290 $F(7734, 13579) = 0.2174$, $p < 2.2 \times 10^{-16}$). In addition, the observed divergence in
291 recombination rate between p_2 and p_1 (which was always expected to be +2 cM/Mb)
292 decreased with increasing levels of gene flow (Figures 3B and 3C, Secondary Contact).
293 This effect would likely result in an increase in false negatives with increasing gene
294 flow (i.e. finding no difference in recombination rate between populations when there is
295 in fact one). This decrease in the observed divergence between populations is again
296 likely the outcome of the population-specific levels of LD becoming coupled/merged at
297 moderate to high levels of gene flow, resulting in the populations exhibiting LD (and
298 hence recombination rate estimates) intermediate to what would be expected in the
299 absence of gene flow.



300 **Figure 3** | The relationship between inferred recombination rate and the migration rate in simulated
 301 populations where recombination rate increases by a factor of two in one subpopulation. (A) Inferred
 302 recombination rates for individual simulations at varying levels of migration. Each plot shows inferred
 303 rates for simulation replicates (transparent lines) of population 1 (red, unchanged recombination) and
 304 population 2 (blue, increased recombination) for a single migration rate. Dashed lines show the expected
 305 inferred value in the absence of gene flow (inferred from $N_e m = 0$). (B) Summarized inferred recombination
 306 rates (y-axis) for each level of migration (x-axis) from the simulations in A. Points are mean values and

307 error bars depict standard deviations (summarized across all generations). Dashed lines show the
308 expected inferred value in the absence of gene flow for each population (i.e. the mean value for $N_e m = 0$).
309 (C) The inferred *difference* in recombination rate between population 1 and population 2 ($p_2 - p_1$) as a
310 function of migration rate. Points and errors bars are as in B.

311 Discussion

312 Accurate estimates of recombination rate are key to understanding the causes and
313 consequences of recombination rate variation in natural populations. With the
314 increasing availability of genome-wide sequencing data, LD-based estimators of
315 recombination rate have become widely used in a large variety of taxa. However, while
316 gene flow is widely known to shape patterns of LD in populations, the effect of gene
317 flow on LD-based estimators of recombination rate remains largely unexplored. Here,
318 use forward-time simulations to show that (1) gene flow can introduce substantial bias
319 into LD-based estimates of genome-wide recombination rate and (2) the nature of this
320 bias depends on the demographic and evolutionary history of the populations in
321 question.

322 Our results here are consistent with theoretical predictions that gene flow between
323 populations can affect LD: increasing in the magnitude and variance of LD at low
324 migration rates as well as reducing LD via the “coupling effect” we observed at higher
325 rates of gene flow. Our study shows how these predictions play out with modern
326 methods and genomic data, and also provides a sense of the magnitude of the potential
327 degree of misestimation - in our case, ranging from 20-50 percentage points in cases of
328 moderate gene flow. For comparison, a recent study of population-level differences in

329 recombination rate in *Drosophila pseudoobscura* revealed genetically based
330 interpopulation differences on the magnitude of approximately 10% measured using
331 replicated linkage maps in each population (Samuk et al., 2020). Using LD-based
332 estimators, an observed difference of this magnitude could be spuriously generated by
333 modest levels of gene flow alone, or missed altogether due to coupling at higher level
334 gene flow is high. In addition, the specific magnitude and direction of the bias
335 introduced by gene flow is difficult to know without precise knowledge of the
336 population/demographic histories of the populations in question. This should give
337 pause to anyone planning on using LD-based methods to infer recombination rate in
338 non-equilibrium populations.

339 One key question is whether there are methods to control for or counteract the
340 increased variance and/or biases in the estimation of recombination rate caused by gene
341 flow. One approach could be to identify and remove introgressed haplotypes from
342 datasets prior to inferring recombination rate, thereby removing migration-associated
343 LD. This would require “pure” samples from the source populations, such that the
344 population of origin could be assigned to haplotype blocks (Dias-Alves et al., 2018).
345 However, this method would only work if gene flow is infrequent enough that coupling
346 (of both LD and allele frequencies) has not occurred. The upward bias and increased
347 variance in recombination rate that occurs as a result of coupling, together with the
348 homogenization of allelic differences between populations at higher levels of gene flow
349 will likely make a “filtering” scheme very difficult (perhaps impossible) to achieve. One
350 approach may be to attempt to jointly estimate a demographic model along with
351 population-specific recombination rates, as has been done with mutation rates (DeWitt

352 et al., 2021). However, given the existing complexity and uncertainty in inferring
353 demographic models, we suspect it may be difficult to disentangle the complex
354 interdependencies between gene flow, population size, and estimates of recombination
355 rate.

356 Together with previous work (Dapper & Payseur, 2018), our results suggest that LD-
357 based estimates of recombination rate need to be interpreted with great caution when
358 studying non-equilibrium populations. Indeed, these methods are likely only
359 appropriate when populations can be assumed to be evolving in the absence of any gene
360 flow, and have reached a reasonable demographic equilibrium. However, it is now
361 widely appreciated that gene flow is ubiquitous in natural populations (Ellstrand &
362 Rieseberg, 2016; Waples & Gaggiotti, 2006). This may mean that many published LD-
363 based estimates of recombination rate are incorrect. Without empirical maps to
364 compare existing LD-based estimates, it is difficult to say just how incorrect. What can
365 be said is that the levels of gene flow required to introduce non-trivial biases into
366 estimates of recombination rate, i.e. $N_e m \sim 1-10$, are not uncommon in natural
367 populations (Slarkin, 1985; Waples & Gaggiotti, 2006). It is also worth noting that it is
368 not the case that two populations being studied have to be exchanging genes themselves
369 (e.g. which would not be the case when studying two reproductively different species), but
370 just that one or more of the populations are exchanging genes with some *other*
371 population (e.g. an unsampled population of the same species).

372 If many LD-based estimates are incorrect, why do published LD-based estimates of
373 recombination rate correlate well with direct estimates, e.g. from genetic maps? (Chan,

374 Jenkins, et al., 2012; Gil McVean & Auton, 2007; Smukowski Heil et al., 2015). There are
375 several considerations. First, the correlations that have been reported are by no means
376 perfect (e.g. \sim Spearman's Rho of 0.6: Smukowski-Heil et al. 2015; $r^2 = 0.37-63$: Chan,
377 Song, et al., 2012) and depend greatly on the genomic scale at which they are measured
378 (Smukowski-Heil et al. 2015). Second, simple correlations between LD-based and
379 empirical estimates cannot detect genome-wide differences in the estimates of
380 recombination rate, such as those due to the coupling effects we observed. Such effects
381 would be visible as differences in the *intercept* of a linear regression, rather than the R^2 ,
382 for example. Finally, the species where these correlations have been examined (humans
383 and *Drosophila melanogaster*) may meet the assumptions of demographic equilibrium
384 more readily (Ochoa & Storey, 2019; Suvorov et al., 2021). While such assumptions may
385 be reasonable for these populations, for which LD-based estimators were originally
386 developed, they are much less likely to hold in many natural populations. Notably, they
387 are likely rarely met in populations that have recently adaptively diverged in the
388 presence of gene flow, which have lately been the subject of increased research interest
389 (Linck & Battey, 2019; Ravinet et al., 2017). The equilibrium assumption is likely not
390 valid in populations in which the recombination rate has recently changed (Brandvain &
391 Wright, 2016), reducing the utility of these estimates for studying the rapid evolution of
392 recombination rates.

393 While we only focused on a single implementation of one type of LD-based estimator of
394 recombination (ρ), it is likely that other population genetic methods will also suffer
395 from the effects we describe here. LD is the “information” used by all estimators, either
396 directly as in methods like LDjump (Hermann et al., 2019) or indirectly as in machine

397 learning methods like ReLERNN (Adrion, Galloway, et al., 2020). That said, in the case
398 of the latter method, it may be possible to overcome some of the issues we've identified
399 if the training datasets were simulated with an accurate demographic model. As such,
400 the distorting effects of gene flow on LD need to be carefully considered when applying
401 any statistical methods for inferring recombination rate approaches. We also stress that
402 our simulations do not suggest that LD-based estimators and their implementations are
403 wrong per se, but rather that the assumptions under which LD-based estimates are
404 biologically accurate are readily violated by levels of gene flow and divergence common
405 seen in natural populations.

406 Conclusion

407 Studying variation in recombination rate is difficult. LD-based methods for inferring
408 recombination rate are attractive in their data requirements, but require strong
409 assumptions to be met. As we have shown here, gene flow readily violates these
410 assumptions and introduces biases and decreases in precision, in a variety of ways that
411 are difficult to identify in a given study population. This is problematic because gene
412 flow is extremely common in natural populations. How should we proceed? Rather than
413 attempt to squeeze blood from the proverbial stone, we believe that the most
414 straightforward solution to the problems we outline here is simply to prioritize the use
415 of direct, empirical methods for measuring of recombination rate. This decision is made
416 hopefully simpler with the increased ease and low cost of creating traditional linkage
417 maps and performing gamete sequencing. That said, LD-based approaches remain
418 important tools for hypothesis generation, and when paired with direct estimates of

419 recombination rate can provide a detailed picture of both the past and present
420 landscape of recombination rates in natural populations.

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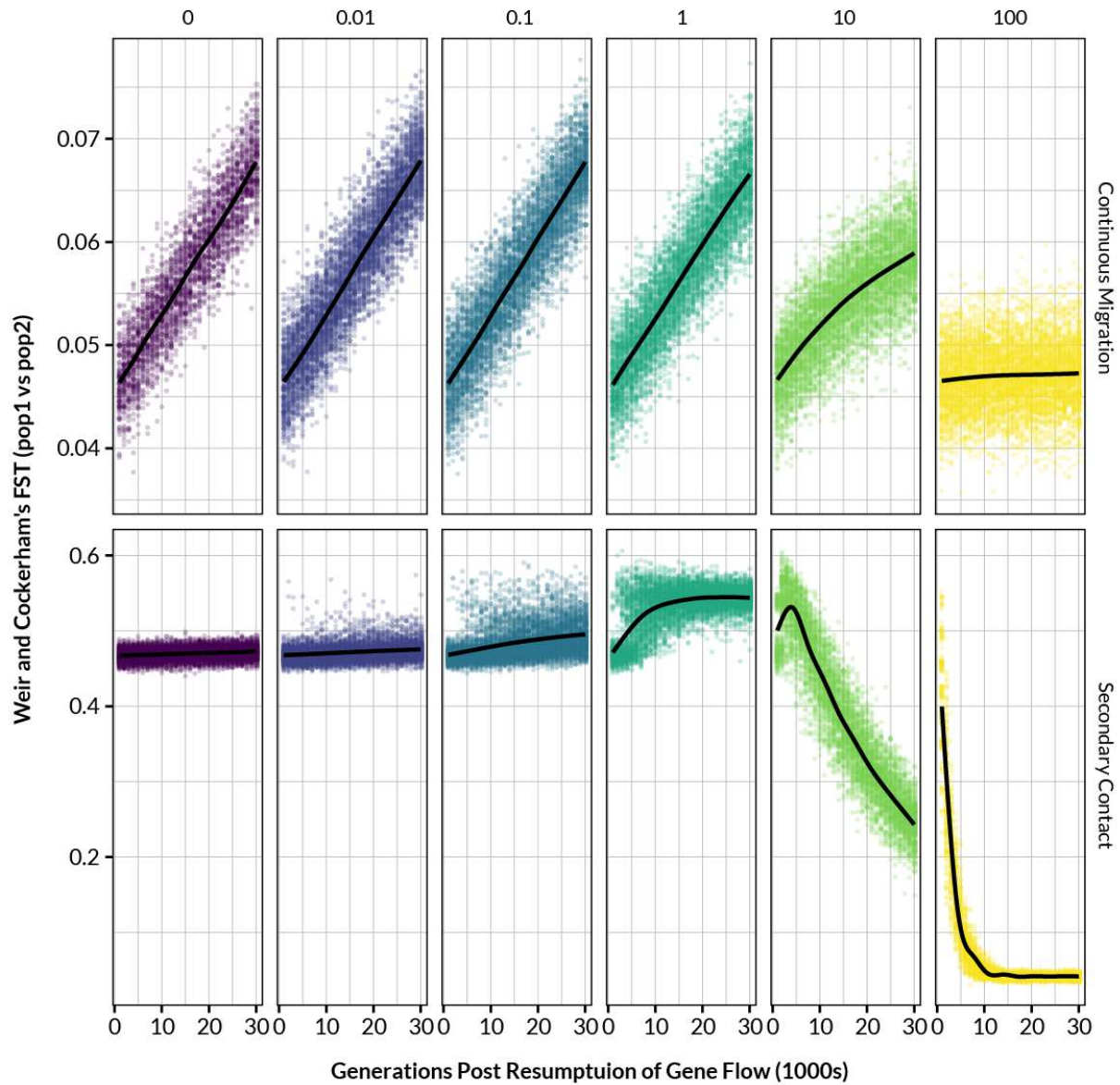
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590

591 Supplemental Material



592
593 **Figure S1** | Weir and Cockerham's FST between simulated populations as a function of time in
594 generations under various combinations of migration rate (columns, $N_e m$) and isolation scenario (rows).
595 Black lines are smoothed LOESS fits. Note the difference in y-axis scales between the rows.